Introduction to resampling methods

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objectives of the course

- to present resampling technics
 - randomization tests
 - cross-validation
 - jackknife
 - bootstrap
- to apply with R



outlines

- 1 Principle and mecanism of the resampling methods
- Randomisation tests
- 3 Cross-validation
- 4 Jackknife
- Bootstrap
- 6 Conclusion





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Resampling in statistics

Description

set of statistical inference methods based on *new* samples drawn from the initial sample

Implementation

- computer simulation of these new samples
- analysing these new data to refine the inference

Classical uses

- estimation/bias reduction of an estimate (jackknife, bootstrap)
- estimation of confidence intervalle without normality assumption (bootstrap)
- exacts tests (permutation tests)
- model validation (cross validation)





History of resampling techniques

1935 randomization tests, Fisher



1948 cross-validation, Kurtz



1958 jackknife, Quenouille-Tukey



1979 *bootstrap*, Efron 🖊







Resampling mecanism

Why it works?

- can we expect an improvement by resampling from the same sample?
- no new information is brought back!
- but it can help to extract useful information from the base sample

the idea

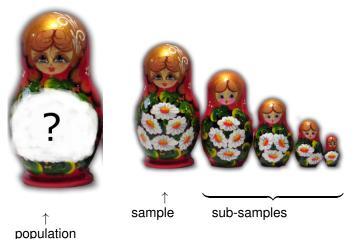
- to consider the sample like the population
- to simulate samples that we could see
- to handle "scale" relationship into the sample





Illustration with Russian dolls

How many flowers is there on the biggest doll?









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randomization tests or permutation tests

goal: testing assumption

 $\begin{cases} H_0: & X \text{ and } Y \text{ are independents} \\ H_1: & X \text{ and } Y \text{ are dependents} \end{cases}$

principle: data are randomly re-assigned so that a p-value is calculated based on the permutated data

permutation tests exhaust all possible outcomes ⇒ exact tests ⇒ exact p-value

randomization tests resampling simulates a large number of possible outcomes ⇒ approximate p-value





example from Fisher: Lady tasting tea

a typical british story

In 1920, R. A. Fisher met a lady who insisted that her tongue was sensitive enough to detect a subtle difference between a cup of tea with the milk being poured first and a cup of tea with the milk being added later. Fisher was skeptical ...

Fisher experiment: he presented 8 cups of tea to this lady

4 cups were 'milk-first' and 4 others were 'tea-first':

















tea or milk first?







example from Fisher: Lady tasting tea

















tea or milk first?

experiment result: the lady well detected the 8 cups

Did the woman really have a super-sensitive tongue?





example from Fisher: Lady tasting tea

















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Reformulation as a statistical test

 H_0 : the order in which milk or tea is poured in a cup and the lady's detection of the order are independent.

 H_1 : the lady can correctly tell the order in which milk or tea is poured in a cup.







example from Fisher: Lady tasting tea

Reformulation as a statistical test

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Reformulation as a statistical test

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probabilities of all the possibilities under H_0

(1,0,1,1,0,0,1,0)	number of match	probability
(1,1,1,1,0,0,0,0)	6	1/nb possibilities
(1,1,1,0,1,0,0,0)	4	1/nb possibilities
:	<u>:</u>	<u>:</u>
(1,0,1,1,0,0,1,0)	8	1/nb possibilities
:	<u>:</u>	:
(0,0,0,0,1,1,1,1)	2	1/nb possibilities





example from Fisher: Lady tasting tea

Reformulation as a statistical test

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test result

The probability of matching the 8 cups under H_0 is 1/(nb possibilities)

number of possibilities
$$\begin{pmatrix} 8\\4 \end{pmatrix} = 70$$

i.e.The probability that the Lady matched by chance the 8 cups is $1/70 \approx 0.014\,$





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exercise

What can we say if the Lady had matched only 6 cups?



example from Fisher: Lady tasting tea

exercise

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some tips to do it with R

- exhausting all possibilities: see function combn
- matching of all the possibilities: loops for or function apply



example from Fisher: Lady tasting tea

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a solution

```
TeaCup <- function(NbCup=8){
    NbMilk=NbCup/2
    ref=rep(0,NbCup)
    ref[sample(NbCup,NbMilk)]=1
    possibilities=combn(1:NbCup,NbMilk)
    score=rep(0,ncol(possibilities))
    for (i in 1:ncol(possibilities)){
        score[i]=sum(ref[possibilities[,i]])*2
    }
    Probabilities=table(score)/choose(NbCup,NbMilk)
    return(Probabilities)
}
```



example: Comparison of two groups

overview

type calories meat 186 181 176 149 184 190 158 139 175 148 152 111 141 153 190 157 131 149 135 132 poultry 129 132 102 106 94 102 87 99 107 113 135 142 86 143 152 146 144



example: Comparison of two groups

Hot-dog data

type	calories
meat	186 181 176 149 184 190 158 139 175 148 152 111
	141 153 190 157 131 149 135 132
poultry	129 132 102 106 94 102 87 99 107 113 135 142 86 143 152 146 144

Hypothesis test

 H_0 : the hot-dog type and calories are independents H_1 : the hot-dog type and calories are dependents

Exercise with R

- Make the test, functions from package coin may be helpful
- What test is it if you use the ranks?





example: test significance of covariate in linear regression

Linear model

$$Y_i = \beta_0 + \beta_1 x_i^1 + \cdots + \beta_p x_i^p + \varepsilon_i \quad i = 1, \ldots, n$$



example: test significance of covariate in linear regression

Linear model

overview

$$Y_i = \beta_0 + \beta_1 x_i^1 + \cdots + \beta_p x_i^p + \varepsilon_i \quad i = 1, \ldots, n$$

hypothesis test

 H_0 : Y and (X_1, \ldots, X_p) are linearly independents

 H_1 : Y and (X_1, \dots, X_p) are linearly dependents



example: test significance of covariate in linear regression

Linear model

$$Y_i = \beta_0 + \beta_1 x_i^1 + \cdots + \beta_p x_i^p + \varepsilon_i \quad i = 1, \ldots, n$$

hypothesis test

 H_0 : Y and $(X_1, ..., X_p)$ are linearly independents H_1 : Y and $(X_1, ..., X_p)$ are linearly dependents

exercise with R

- simulate data according to the model
- proceed a permutation test with the function Imp from the package ImPerm







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Cross-validation

goal

to estimate how accurately a predictive model will perform in practice

principle

- splits the dataset into training and validation data
- fit model on training data
- assess predictive accuracy on validation data (RMSE)

some types of cross-validation

random subsampling (randomly splits the data into training and validation data) $\times s$

K fold randomly splits the data into K subsamples (K-1) for training and 1 for validation $\times K$

leave-one-out (N-1) obs for training and 1 obs for validation)×N





Example: K-fold

- data: $(x_i, y_i)_{i=1,...,100}$
- 5 folds: $d_1 = (x_i, y_i)_{i=1,...,20}, d_2 = (x_i, y_i)_{i=21,...,40}, d_3 = (x_i, y_i)_{i=41,...,60}, d_4 = (x_i, y_i)_{i=61,...,80}$ and $d_5 = (x_i, y_i)_{i=81,...,100}$

CV for model \mathcal{M} : $\hat{y} = f(x, \theta)$

- assess prediction for subsample d₁
 - **1** training: $\theta^1 = \arg \min_{\theta} \sum_{i=21}^{100} (f(x_i, \theta) y_i)^2$
 - 2 predicting: $\hat{y}_i = f(x_i, \theta^1)$ for i = 1, ..., 20
 - **3** validation: $MSE_1 = \sum_{i=1}^{20} (\hat{y}_i y_i)^2 / 20$
- idem for $d_2, \ldots, d_5 \rightarrow \textit{MSE}_2, \ldots, \textit{MSE}_5$
- $CV = mean(MSE_1, MSE_2, MSE_3, MSE_4, MSE_5)$







excercise with R

linear regression

• simulate data according to the model $y = \beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k + \varepsilon$

compute CV for several covariate combination (use function cv.glm form package boot



excercise with R

linear regression

• simulate data according to the model $y = \beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k + \varepsilon$

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generalised linear model

- simulate data according to the model $y \sim \mathcal{P}(\exp(\beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k))$
- compute CV for several covariate combination (use function cv.glm form package boot





excercise with R

linear regression

simulate data according to the model

$$y = \beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k + \varepsilon$$

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generalised linear model

- simulate data according to the model $y \sim \mathcal{P}(\exp(\beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k))$
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any model

- see function crossval from package boostrap
- make your own function



Cross-validation caracteristics

generalities

- easy to implement, the prediction method need only be available as a "black box"
- provides ECV a nearly unbiased estimate of the expectation of the fitting criterion
- can be very slow since the training must be carried out repeatedly, methods to speed up exists but . . .

limitations

- validation set and test set must drawn from the same population
- be careful with dynamical systems and stratified population
- instability for small sample
- ullet variance of CV can be large o to take into account for model selection





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the Jackknife

overview

- proposed to reduce the bias of an estimator (delete if bias = a/n)
- generally used to estimate the variance or covariance
- useless for unbiased estimators (mean)
- close to leave one out method
- only adapted for estimators which are smooth function of the observation

principle

- let a sample $x_1, \ldots, x_n \sim \mathcal{L}(\theta)$
- let T an estimator of θ computed with x_1, \ldots, x_n
- for i = 1, ..., n: T_{-i} is the estimator computed with $x_1, ..., x_{i-1}, x_{i+1}, ..., x_n$
- the Jackknife estimators: $T_J = T \frac{n-1}{n} \sum_{i=1}^{n} (T_{-i} T)$
- bias of the estimator: $-\frac{n-1}{n}\sum_{i=1}^{n}(T_{-i}-T)$



excercise: Jackknife estimator for the variance

estimators of variance

Let $x_1, \ldots, x_n \sim \mathcal{N}(\mu, \sigma^2)$, the two usual estimators of σ^2 :

$$S^2 = \frac{1}{n} \sum_{i=1}^{n} (x_i - \bar{x})^2$$
 and $\tilde{S}^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2$

their expectations:

$$E[S^2] = \frac{n-1}{n}\sigma^2$$
 and $E[\tilde{S}^2] = \sigma^2$





excercise: Jackknife estimator for the variance

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their expectations:

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 and $E[\tilde{S}^2] = \sigma^2$

exercise with R

- simulate a sample
- compute the jackknife estimator of S^2 using the function *jackknife* from package *bootstrap*
- compare with the estimator \tilde{S}^2





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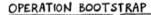
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The bootstrap















The bootstrap

goal

- estimate confidence interval of a parameter with or with assumption on data distribution
- estimate bias of an estimator

principle of the nonparametric bootstrap

- starting point: x_1, \ldots, x_n *i.i.d* $\sim \mathcal{L}(\theta)$
- repeat for $k = 1, \dots, K$
 - **1** sample with replacement x_1^*, \ldots, x_n^* among x_1, \ldots, x_n
 - **2** compute θ_k^* the estimates of θ with x_1^*, \ldots, x_n^*
- $\theta_1^*, \dots, \theta_K^* \leadsto \text{inference on } \theta$

many variants

- parametric bootstrap
- smooth bootstrap . . .



The bootstrap: generalities

- In Jackknife, the number of resamples is confined by the number of observations (n-1).
- In bootstrap, the original sample could be duplicated as many times as the computing resources allow, and then this expanded sample is treated as a virtual population.
- Unlike cross validation and Jackknife, the bootstrap employs sampling with replacement
- In cross-validation and Jackknife, the n in the sub-sample is smaller than that in the original sample
- In bootstrap every resample has the same number of observations as the original sample.
- Thus, the bootstrap method has the advantage of modeling the impacts of the actual sample size



example: median estimation based on bootstrap

sample: 10, 15, 17, 8 and 11 \rightarrow median estimate = 11

resampling	median
8 11 15 10 8	10
10 15 10 10 11	10
15 17 10 10 17	15
11 15 10 10 17	11
17 17 15 17 10	15
10 15 10 17 10	10
10 8 8 8 15	8

median estimate by bootstrap: mean(10,10,15,11,15,10,8)=11.29 sd of median estimate by bootstrap: sd(10,10,15,11,15,10,8)=2.69





example: median estimation based on bootstrap

sample: 10, 15, 17, 8 and 11 \rightarrow median estimate = 11

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10 15 10 17 10	10
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median estimate by bootstrap: mean(10,10,15,11,15,10,8)=11.29 sd of median estimate by bootstrap: sd(10,10,15,11,15,10,8)=2.69

exercise with R

estimate the confidence interval using the function *bootstrap* from package *bootstrap*





example: parametric bootstrap

principle of the parametric bootstrap

- starting point: x_1, \ldots, x_n i.i.d $\sim \mathcal{L}(\theta)$ where \mathcal{L} known
- compute $\hat{\theta}$ the mle of θ
- repeat for k = 1, ..., K
 - igotimes sample x_1^*,\ldots,x_n^* i.i.d $\sim \mathcal{L}(\hat{\theta})$
 - **2** compute θ_k^* the mle of θ with x_1^*, \ldots, x_n^*
- $\theta_1^*, \dots, \theta_K^* \rightsquigarrow \text{inference on } \theta$



example: parametric bootstrap

principle of the parametric bootstrap

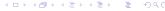
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Exercise with R: the exponential distribution

- simulate a sample from $Exp(\lambda)$ with $\lambda = 5$
- estimate λ using both parametric bootstrap and nonparametric bootstrap (function boot from package boot)







example: resampling residual

principle of the resampling residual

- starting point: $(x_1, y_1), \dots, (x_n, y_n)$ and a model $f(\cdot, \theta)$
- calibrate the model $\rightarrow \hat{\theta}$: $\hat{y}_i = f(x_i, \hat{\theta}) \rightsquigarrow \varepsilon_i = \hat{y}_i y_i$
- repeat for $k = 1, \dots, K$
 - **1** $\forall i, y_i^* = y_i + \tilde{\varepsilon}$ with ε sampled from $\varepsilon_1, \ldots, \varepsilon_n$
 - 2 calibrate the model with $(x_1, y_1^*), \ldots, (x_n, y_n^*) \rightsquigarrow \theta_k^*$
- $\theta_1^*, \dots, \theta_K^* \leadsto \text{inference on } \theta$

exercise with R: the case of the linear model

- simulate data according to a linear model
- estimate regression parameters using the function *lm.boot* from the package *simpleboot*







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Rationale of supporting resampling

Empirical empirical-based resampling do not require assumptions on the sample or the population.

Clarity resampling is clean and simple. High mathematical background is not required to comprehend it

Small sample size Distributional assumptions requierd by classical procedures are usually met by a large sample size.

Bootstrapping could treat a small sample as the virtual population to "generate" more observations

Non-random sample Resampling is valid for any kind of data, including random and non-random data.

Large sample size Given a very large sample size, one can reject virtually any null hypothesis \rightarrow divide the sample into subsets, and then apply a simple or double cross-validation.

Replications Repeated experiments in resampling such as cross-validation and bootstrap can be used as internal replications.

Criticisms of resampling

Assumption "You're trying to get something for nothing". Every theory and procedure is built on certain assumptions and requires a leap of faith to some degree. Indeed, the classical statistics framework requires more assumptions than resampling does

Generalization resampling is based on one sample and therefore the generalization cannot go beyond that particular sample.

Bias confidence intervals obtained by simple bootstrapping are always biased though the bias decreases with sample size. (for normal case the bias in is at least n/(n-1))

Accuracy for small sample resampling may be less accurate than conventional parametric methods. Not very convincing argument because today computers are very powerful.

pros and cons in both traditional and resampling methods carry certain valid points. \rightarrow the appropriateness of the methodology highly depends on the situation

R packages for resampling methods

boot quite a wide variety of bootstrapping tricks.

bootstrap relatively simple functions for bootstrapping and related techniques.

coin permutation tests

Design includes bootcov for bootstrapping the covariance of estimated regression parameters and validate for testing the quality of fitted models by cross validation or bootstrapping.

MChtest Monte Carlo hypothesis tests: tests using some form of resampling.

meboot a method of bootstrapping a time series.

permtest a function for permutation tests of microarray data.

resper for doing restricted permutations.

scaleboot produces approximately unbiased hypothesis tests via bootstrapping.

simpleboot performs bootstraps in simple situations: one and two samples, and linear regression.